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PERFORMANCE ANALYSIS OF DISCRETE TIME BULK SERVICE QUEUING MODEL NB $(L,K)/Geo/1$

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ABSTRACT

The present paper is an attempt to explore a discrete – time queueing models which is suitable for the performance evaluation of Asynchronous Transfer Mode (ATM) multiplexers and switches. In these models, the time axis is divided into fixed-length slots and the service of a customer must start and end of slot boundaries. Most analytical studies of discrete-time queues assume constant service times equal to one slot, an infinite buffer capacity and/or uncorrelated arrival process. A discrete time bulk service rule (L, K) , the inter arrival time are assumed to be independent and identically distributed random variable follows negative binomial distribution. The time arrival between the consecutive arrivals of customers is described in terms of a discrete probability mass function (p.m.f) $g(x)$; $g(x)$ is the probability of having an interval of an integer number of x time units between the arrival customer number n and customer number $n+1$. The service time of customers n is given in terms of a discrete p.m.f $s(k)$ follows geometric distribution. The arrival and service are independent and identically distributed (i.i.d). The arrival of customer in First In First Out (FIFO) order. One server transports packets in batches of size minimum L packets and maximum K . The steady state analysis for the considered queueing system is presented and the generating functions of the number of customers in the system are obtained. We also obtained the closed form of expressions for some performance measures of the system.

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INTRODUCTION

In the study of discrete-time queue of packets and the distribution of system occupancy just before and after the departure epochs follows the bulk service rule (L, K) . A fundamental motive for studying discrete-time queues is that they become more appropriate than their continuous-time counterparts for analysing computer and telecommunication systems. Since these systems are more digital (machine cycle times, bits and packets, etc.) than analogue. In view of this, there have an increasingly important owing to their applications in the study of many computer and communication systems such as asynchronous transfer mode (ATM), multiplexers in Broadband Integrated Service Digital Networks (BISDN's), Circuit-Switched Time Division Multiple Access (TDMA), Carrier Sense Multiple Access (CSMA) protocols and traffic concentrators in which the time axis is divided into slots. The packets arrive one by one inter-arrival time following negative binomial distribution, with the probability of success of arrival in a slot is 'p' and not for arrival is 'q'. If an arriving packet finds the server busy, it joins the queue in First In and First Out (FIFO) order. The service times of batches are assumed to constitute a set of i.i.d positive random variable with a geometric distribution with probability p_1 , and the complementary is q_1 .

Chaudhry and Templeton [1] have presented extensive discussions of bulk service systems that operate according to a rule admitting non-accessible batch. Mathias [6] studies a inter-departure time distribution for batches and studies correlation between inter-departure time and batch sizes. Mathias and Alexander [5] provide an discrete-time analysis in the performance evaluation of manufacturing systems. Sivasamy and Elangovan [11] on the other hand, the batch, which entries service queue with accessible and non-accessible batches. Goswami, *et al.*, [3] discrete-time bulk-service queues with accessible and non-accessible batches. Sivasamy and Pukazhenthhi [12] have carried out and analyzed the discrete time bulk service queue for the accessible batch with the arrivals time being geometrical distribution and services time being negative binomial distribution. Vijaya Laxmi Pikala, *et al.*, [13] have studied the discrete-time renewal input bulk service queue with changeover time. Pukazhenthhi and Ezhilvanan [10] discussed the discrete time queue length distribution with a bulk service rule. Daniel Wei-Chung Miao, *et al.*, [2] computational analysis of a markovian queueing system with geometric mean-reverting arrival process. Pukazhenthhi and Ezhilvanan [8] the analysis of

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discrete time queues with single server using correlated times. Pukazhenthil and Ezhilvanan [7] analysis of discrete time queue of packets with G/G/1 queueing model. Pukazhenthil and Ezhilvanan [9] recently used a vacation queueing model with bulk service rule. In this present study, this paper focus on the analysis of discrete time queue packets with bulk service rule (L,K). Customers arrive according to negative binomial arrival process with parameter p , the probability that an arrival occurs in a slot and the service time is assumed to follow geometric distribution with parameter p_1 . We consider on (L,K) policy bulk service queue with accessible and non accessible batches under the customer's choice. In this model the time axis is divided into fixed-length of contiguous intervals, referred to as slots and it is assumed that the inter-arrival and the service times are integral multiples of slot durations. It is always assumed that service time can be started and completed only at slot boundaries and that their durations are integral multiples of slot durations. Customers arrive to the system in according with a negative binomial process and are accommodated in the buffer. The numbers of customers entering the system during the consecutive slots are assumed to be (i.i.d) non negative discrete random variables in a discrete case. A batch of packets of size j ($L \leq j \leq K$) starts service at the beginning of a slot, and may only end service just before the end of a slot. The probability of a batch being transported by the server at the end of a slot is probability ' p_1 ' and the probability of the batch not being transported is ' $1 - p_1$ '. Service time of a batch is independent of the number of packets in a batch and simultaneous occurrence of both arrival and departure in a single slot is ruled out. All the random variables in the analysis are nonnegative and of integer valued. It is also assumed that the service times and the arrival process are mutually independent. The customers are served in batches of variables capacity, the maximum service capacity being a finite integer value. The steady state behaviours of the models are studied and explicit expressions for the steady state probabilities and the expected queue length are obtained.

The rest of the paper is organized as follows: In section 2, presents the description of the model, i.e, arrival distribution and service distribution. Section 3, deals with the steady state distribution of system occupancy at departure epochs, the joint distribution of inter-departure time and number of packets in a batch. In section 4 presents numerical study for various performance measures of the steady state distribution of the batch server at departure instants. A brief conclusion is presented in section 5.

Description of the Model

Let assume that before customers' arrival follow the negative binomial distribution, and the customer's service follow the geometric distribution, with the queue of packets in an infinite capacity, on the bulk server rule of minimum size L and maximum size K. The packets are independent and identically distributed, $g(x)$; (i.i.d) the random variables.

Inter Arrival-time Distribution

Let $p(x)$ be the number of customers those arrive during slot x . In this model it is assumed that the inter-arrival ' p ' of customers are independent and negative binomial distribution $\{a(x; \alpha, p) = Pr (A_n = x) \}$ and the arrival time are independent of the number of slots in the batch.

$$a(x; \alpha, p) = \binom{x-1}{\alpha-1} p^\alpha (1-p)^{x-\alpha}; \{x = \alpha, \alpha + 1, \alpha + 2 \dots\} \quad \dots (1)$$

Hence, the mean arrival times $\left(\frac{\alpha}{p}\right)$ and variance of the arrival times are $\left(\frac{\alpha}{p^2}\right)$

$$\text{i.e, } E(A) = \alpha/p, \quad E(A^2) = \alpha(\alpha + q)/p^2 \quad \text{and} \quad \text{Var} (A) = \frac{\alpha}{p^2} \quad \dots (2)$$

The mean number of customers in the service time distribution

The service time follows geometric distribution and service of packets are independent and identically distributed random variable of $s(x)$ with probability p_1 . The service time for the n th batch is of length B_n where B_n are random variables that follows geometric distribution with (p.m.f) $b(k)$, Service times are independent of the number of slots in the batch $b(k) = Pr (B_n = k); \{k = 0,1,2, \dots\}$

$$b(k) = p_1(1 - p_1)^{k-1}; k = 0,1,2 \dots \quad \dots (3)$$

where ' p_1 ' probability of service time,

B_n is the number of slots required to complete a batch service at the k^{th} success in a sequence of independent Bernoulli trials with probability p_1 for success and $q_1 = 1 - p_1$ is the probability of failure. The mean, $E (B^2)$ and $\text{var} (B)$ are

$$\text{i.e, } E(B) = 1/p_1, \quad E(B^2) = (2 - p_1)/p_1^2 \quad \text{and} \quad \text{var}(B) = q_1/p_1^2 \quad \dots (4)$$

Utilization Factor

Utilization is the proportion of the system resources which is used by the traffic which arrives at it (Negative binomial arrivals and Geometric server), then it is given by the mean arrival rate over the mean service rate that is

$$\rho = \frac{E(B)}{KE(A)} = \frac{p}{Kp_1} < 1 \quad \dots (5)$$

Distribution of System Occupancy at Departure Epochs

Let X_n be the number of packets which has $0,1,2,\dots,\infty$ accumulated in the system (queue + service) just after the server has left with the n^{th} batch. The steady state distribution $\{x(k) = \lim_{n \rightarrow \infty} x_n(k) = \lim_{n \rightarrow \infty} \Pr(X_n = k) : k = 0,1,2,\dots,\infty\}$ of system occupancy at departure epochs is derived in this section using the embedded Markov Chain (MC) technique.

Let v_n be random variable denote the number of packets that reaching the system during the n^{th} service then the distribution $\{v_n(k) : k = 0,1,2,\dots\}$ can be derived as follows:

$$v(k) = \sum_{x=k+\alpha}^{\infty} \binom{x-\alpha}{k} \binom{x-1}{\alpha-1} p^\alpha (1-p)^{x-\alpha} p_1^k (1-p_1)^{x-\alpha-k} \quad \dots (6)$$

$$k = 0,1,2,\dots,\infty$$

Put $r = x - \alpha - k$ which gives $r + k = x - \alpha$ in the above equation. Further as $r + k + \alpha = x$ gives $x = k + \alpha$, when $r = 0$ and $x = \infty$ implies $r = \infty$. Therefore the above equation can be rewritten as

$$v(k) = \sum_{x=k+\alpha}^{\infty} \binom{x-\alpha}{k} \{p^\alpha (1-p)^{x-\alpha}\} \{p_1^k (1-p_1)^{x-\alpha-k}\} \binom{r+k+\alpha-1}{\alpha-1}$$

$$v(k) = \{p^\alpha p_1^k (1-p_1)^k\} \sum_{r=0}^{\infty} \binom{r+k}{k} (1-p)^r \binom{r+k+\alpha-1}{\alpha-1} (1-p_1)^r$$

$$v(k) = \{p^\alpha p_1^k (1-p_1)^k\} \sum_{r=0}^{\infty} \frac{(k+\alpha-1)! (r+k+\alpha-1)!}{k! r! (\alpha-1)! (k+\alpha-1)!} (1-p)^r (1-p_1)^r$$

$$v(k) = \{p^\alpha p_1^k (1-p_1)^k\} \sum_{r=0}^{\infty} \frac{(r+k)!}{r! k!} \frac{(r+k+\alpha-1)!}{(r+k)! (\alpha-1)!} (1-p)^r (1-p_1)^r$$

$$v(k) = \{p^\alpha p_1^k (1-p_1)^k\} \frac{(k-\alpha-1)!}{k! (\alpha-1)!} \sum_{r=0}^{\infty} \binom{r+k+\alpha-1}{\alpha-1} (1-p)^r (1-p_1)^r$$

$$v(k) = \binom{k+\alpha-1}{\alpha-1} \frac{p^\alpha p_1^k (1-p_1)^k}{(1-(1-p)(1-p_1))^{k+\alpha}}$$

where $\sum_{r=0}^{\infty} \binom{r+k}{k} d^r = 1/(1-d)^{k+1}$

$$v(k) = \binom{k+\alpha-1}{\alpha-1} \frac{p^\alpha p_1^k (1-p_1)^k}{(1-1+p_1+p-pp_1)^{k+\alpha}}$$

$$v(k) = \binom{k+\alpha-1}{\alpha-1} \left(\frac{p}{p+p_1-pp_1}\right)^\alpha \left(\frac{p_1(1-p)}{p+p_1-pp_1}\right)^k$$

Hence

$$V(k) = \binom{k+\alpha-1}{\alpha-1} \beta^\alpha (1-\beta)^k \quad \dots (7)$$

Where

$$\beta = \frac{p_1}{p+p_1-pp_1}$$

Also,

$$E(V) = \frac{\alpha(1-\beta)}{\beta}$$

$$Var(V) = \frac{\alpha(1-\beta)}{\beta^2}$$

The sequence $\{X_n\}$ of random variables can be shown to form a MC on the discrete state space $\{0,1,2,\dots,\infty\}$ with the following one step transition probability matrix $\mathbf{P} = (p_{ij})$ where

$$p_{ij} = \begin{cases} \sum_{r=0}^{K-L} v(r) & ; 0 \leq i \leq L-1 \text{ and } j = 0 \\ v(K-L+j) & ; 0 \leq i \leq L-1 \text{ and } j \geq 1 \\ \sum_{r=0}^{K-i} v(r) & ; L \leq i \leq K \text{ and } j = 0 \\ v(K-i+j) & ; L \leq i \leq K \text{ and } j \geq 1 \\ v(j+K-i) & ; i \geq (K+1) \text{ and } j \geq (i-K) \\ 0 & ; \text{otherwise} \end{cases} \dots (8)$$

The unknown probability (row) vector $X_n = (x_0, x_1, x_2, \dots)$ can now be obtained solving the following system of equations.

$$P = X_n, X_n e = 1, \dots (9)$$

Where 'e' denotes the row vector of unities. A number of numerical methods could be suggested to solve the system of equations (9). For example an algorithm for solving the system of equations is given by Latouche and Ramaswami [4]. Here an upper bound N and 'i' and 'j' of unit step probability function p_{ij} has been selected so that the p_{ij} values are very small for all $i, j \geq N$ and they could be ignored. Thus $\sum_{j=0}^N p_{ij} = 1$ and $p_{iN} = 1 - \sum_{j=0}^{N-1} p_{ij}$ for all $0 \leq i \leq N$ and $P = (p_{ij})$ is a square matrix of order(N + 1).

The expected average system length L_s of the model is calculated as

$$L_s = \sum_{n=0}^{\infty} n x(n) \dots (10)$$

The expected average queue length L_q of the model is calculated as applying equation

$$L_q = \sum_{n=0}^{L-1} n x(n) + \sum_{n=K}^{\infty} (n - K) x(n) \dots (11)$$

The average system length L_s and average queue length L_q can be obtained, for certain parameter value are, presented in the following section.

NUMERICAL RESULTS

Some numerical results have been obtained and presented in the form of a table which is self-explanatory.

With the various probability value of p, for some fixed values L, K, p_1 , and α , the average system length and average queue length are given in the table-1 and also as a graph.

Table 1 Numerical values

$p_1=0.10$	L=3	K=12	$\alpha = 1$
P	L_s	L_q	
0.98	12.5309	6.975035	
0.97	11.3335	6.507471	
0.96	11.3382	6.070672	
0.95	10.7910	5.664622	
0.94	10.7910	5.284090	
0.93	9.7805	4.930805	
0.92	9.3145	4.600096	
0.91	8.8722	4.291174	
0.90	8.8524	4.291174	
0.89	8.0544	3.734058	
0.88	7.6752	3.482308	
0.87	7.3144	3.246895	

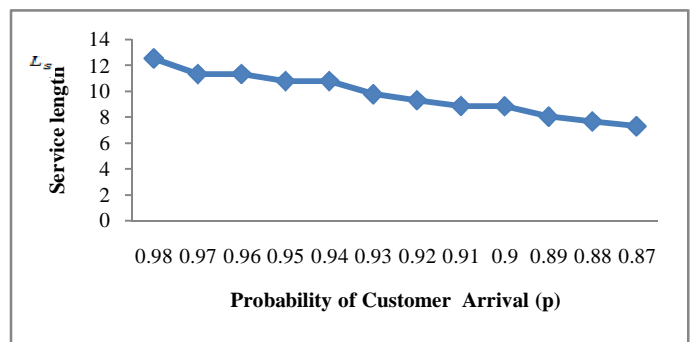


Fig 1 a

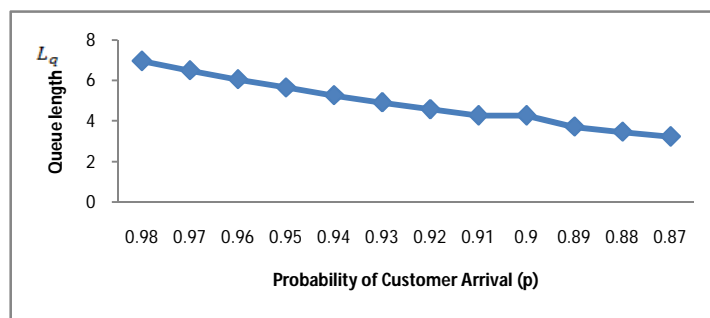


Fig 1 b

From fig-1(a) and fig-1(b), we can see that the expected system length L_s and expected queue length L_q decreases with the rate of probability of arrival decreases. And the larger interruption probability p is, the expected system length L_s and expected queue length L_q becomes larger which is in agreement with practical case.

With the various probability value of p_1 , for fixed the values L, K, p , and α , the average system length and average queue length are given in the table-2 and also as a graph.

Table 2 Numerical values

$p=0.98$	$L=3$	$K=12$	$\alpha = 1$
p_1	L_s		L_q
0.10	12.5309		6.975035
0.12	5.1885		1.958664
0.14	2.7194		0.748689
0.16	1.5895		0.346206
0.18	0.9886		0.185701
0.20	0.6392		0.112171
0.22	0.4242		0.074363
0.24	0.2863		0.052489
0.26	0.1953		0.038437
0.28	0.1343		0.028927
0.30	0.0929		0.022161
0.32	0.0641		0.016757

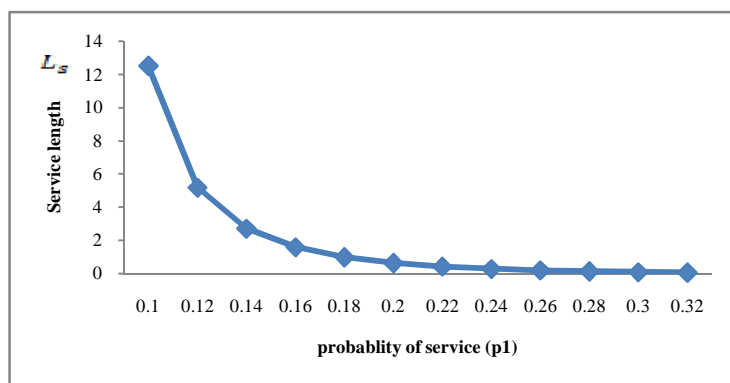


Fig 2 a

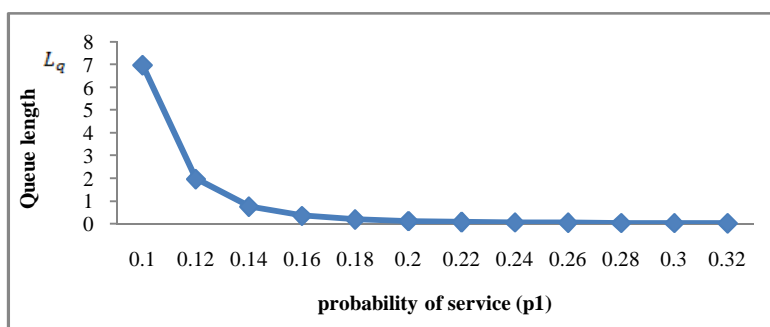


Fig 2 b

Fig-2(a) and Fig-2(b) provide the expected queue length and expected system length, respectively, with the service rate different probability p_1 . It is obvious that L_s and L_q decreases evidently with the rate p increases and the larger probability p_1 is, the smaller L_s become smaller. Therefore under the bulk service policy (L,K) if we want to develop a better service, we can consider (L,K) policy which utilizes the server and decreases the waiting jobs effectively.

CONCLUSION

In this paper, we propose a new $NB(L,K)/G/1$ queue with bulk service rule (L,K) . The system is more complicated and generalizes many discrete time queues about bulk service rule. Using matrix geometric method, we obtain the steady-state distributions and some performance measure. We also perform some numerical examples to study the effect of various parameters on the expected queue length and expected system length. For future research, one can consider for some model when the bulk rule follow those type distributions.

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